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A sequential test with three possible decisions for comparing two unknown probabilities, based on groups of observations

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1. Introduction.

We consider two series of independent trials, e.g. two processes A and B, each trial resulting in a success or a failure with probabilities $p, 1-p$ and $p', 1-p'$ for the two processes respectively.

A sequential test with two possible decisions for the comparison of p and p' , developed by WALD [6], may be used if the trials are executed in pairs, each pair consisting of one trial for each process.

For groups of trials of both processes a sequential test with two possible decisions for comparing p and p' has been described in [8]. This test is carried out as follows.

Suppose the group of trials, constituting the i -th stage of the test, consists of n_i trials for process A and n'_i for process B. If the numbers of successes are a_i and b_i ¹⁾ respectively, a_i and b_i both possess a binomial probability distribution with parameters n_i, p and n'_i, p' respectively.

The following transformation is then used: if n' possesses a binomial probability distribution with parameters n and p the random variable

$$(1.1) \quad \begin{cases} y = 2 \arcsin \sqrt{\frac{n'}{n}} & 0 < n' < n \\ y = \sqrt{\frac{2}{n}} & n' = 0 \\ y = \pi - \sqrt{\frac{2}{n}} & n' = n \end{cases}$$

is, for large n , approximately normally distributed with mean

$$(1.2) \quad \mu = 2 \arcsin \sqrt{p}$$

and variance

$$(1.3) \quad \sigma^2 = \frac{1}{n} + \frac{1}{n^2} \cdot 2)$$

The transformation (1.1) is applied to a_i and b_i ; after this transformation, these random variables will be denoted by \underline{u}_i and \underline{v}_i respectively.

1) Random variables are denoted by underlined symbols; the same symbols, not underlined, are used to denote values assumed by these random variables.

2) Tables of y (in radians) and σ^2 are given in [8] for $n = 10(1)50$ and $0 \leq n' \leq n$.

The sequential test of WALD [6] with two possible decisions for the mean of a normal distribution with known variance is then applied to

$$x_i = u_i - v_i \quad (i=1,2,\dots).$$

Both abovementioned tests are tests with two possible decisions, i.e. the tests result in one of the decisions $p > p'$ or $p < p'$.

The sequential test for comparing p and p' for the case of pairs of trials may be generalized to a test with three possible decisions, i.e. a test resulting in one of the decisions $p > p'$, $p < p'$ or $p \approx p'$, by means of a test developed by DE BOER [2]. This case will not be considered here.

In this paper the abovementioned test for comparing p and p' for groups of trials will be generalized by means of a sequential test for the mean of a normal distribution with known variance developed by SOBEL and WALD [7].

The sequential test of WALD with two and the test of SOBEL and WALD with three possible decisions for the mean of a normal distribution with known variance will be described first.

2. Sequential test for the mean of a normal distribution with known variance.

2.1. Two possible decisions.

For the case that the successive observations x_1, x_2, \dots are independent observations of one random variable x , possessing a normal probability distribution with mean μ and known variance σ^2 , WALD's sequential test with two possible decisions for μ has been described in [6] (p. 117-124). This test will be described here for the case that the variance is not constant. This results in a small change in WALD's test; the proof of the validity of this test follows at once from WALD's own proofs.

For the test a value μ_0 of μ must be chosen, the two possible decisions being: $\mu < \mu_0$ and $\mu > \mu_0$, where we may substitute \leq resp. \geq for $<$ resp. $>$.

Furthermore two values μ_1 and μ_2 must be chosen with

$$\mu_1 < \mu_0 < \mu_2$$

such that the decision $\mu > \mu_0$ is considered an incorrect decision if $\mu \leq \mu_1$ and the decision $\mu < \mu_0$ is considered incorrect if $\mu \geq \mu_2$; for values of μ between μ_1 and μ_2 it is not important which decision is taken.

The concepts "correct" and "incorrect decision" are thus

defined as follows:

Table I
Correct and incorrect decisions

value of μ	correct	incorrect
	decision	
$\mu \leq \mu_1$	$\mu < \mu_0$	$\mu > \mu_0$
$\mu \geq \mu_2$	$\mu > \mu_0$	$\mu < \mu_0$
$\mu_1 < \mu < \mu_2$	$\begin{cases} \mu < \mu_0 & \text{and} \\ \mu > \mu_0 \end{cases}$	—

The interval (μ_1, μ_2) is called the indifference region.

If:

α = the probability of acceptance of $\mu > \mu_0$ if $\mu = \mu_1$,

β = the probability of acceptance of $\mu < \mu_0$ if $\mu = \mu_2$,

and if α and β are chosen both $< \frac{1}{2}$ the probability of an incorrect decision is $\leq \alpha$ for $\mu \leq \mu_1$ and $\leq \beta$ for $\mu \geq \mu_2$.

The value μ_0 is of no further importance for the performance of the test.

The test is carried out as follows (σ_i^2 is the known variance of x_i):

Additional observations are taken as long as:

$$(2.1) \quad \frac{\ln B}{\mu_2 - \mu_1} < \sum_{i=1}^n \frac{x_i - \frac{\mu_1 + \mu_2}{2}}{\sigma_i^2} < \frac{\ln A}{\mu_2 - \mu_1},$$

where

$$A = \frac{1 - \beta}{\alpha} > 1$$

$$B = \frac{\beta}{1 - \alpha} < 1.$$

The test is terminated as soon as (2.1) does not hold and the decision $\mu > \mu_0$ is then taken if

$$\sum_{i=1}^n \frac{x_i - \frac{\mu_1 + \mu_2}{2}}{\sigma_i^2} \geq \frac{\ln A}{\mu_2 - \mu_1}$$

and the decision $\mu < \mu_0$ if

$$\sum_{i=1}^n \frac{x_i - \frac{\mu_1 + \mu_2}{2}}{\sigma_i^2} \leq \frac{\ln B}{\mu_2 - \mu_1}.$$

If the random variables α_i all have the same variance σ^2 the test may be carried out graphically, as indicated by WALD [6] (p. 118-121).

2.2. Three possible decisions.

The sequential test with three possible decisions for the mean μ of a normal distribution with known variance, developed by SOBEL and WALD has been described in [7] for the case that the variance of α_i is a constant. This restriction is again dropped here.

For the test two values μ_0 and μ'_0 and four values μ_1, μ_2, μ_3 and μ_4 must be chosen such that

$$\mu_1 < \mu_0 < \mu_2 < \mu_3 < \mu'_0 < \mu_4,$$

the three possible decisions being:

1. $\mu < \mu_0$
2. $\mu > \mu'_0$
3. $\mu_0 \leq \mu \leq \mu'_0$.

The intervals (μ_1, μ_2) and (μ_3, μ_4) are the indifference regions.

The concepts "correct" and "incorrect decision" are defined as follows:

Table II
Correct and incorrect decisions

value of μ	correct decision	incorrect
$\mu \leq \mu_1$	$\mu < \mu_0$	$\left\{ \begin{array}{l} \mu_0 \leq \mu \leq \mu'_0 \\ \mu > \mu'_0 \end{array} \right.$
$\mu_1 < \mu < \mu_2$	$\left\{ \begin{array}{l} \mu < \mu_0 \\ \mu_0 \leq \mu \leq \mu'_0 \end{array} \right.$	$\mu > \mu'_0$
$\mu_2 \leq \mu \leq \mu_3$	$\mu_0 \leq \mu \leq \mu'_0$	$\left\{ \begin{array}{l} \mu > \mu'_0 \\ \mu < \mu_0 \end{array} \right.$
$\mu_3 < \mu < \mu_4$	$\left\{ \begin{array}{l} \mu_0 \leq \mu \leq \mu'_0 \\ \mu > \mu'_0 \end{array} \right.$	$\mu < \mu_0$
$\mu \geq \mu_4$	$\mu > \mu'_0$	$\left\{ \begin{array}{l} \mu < \mu_0 \\ \mu_0 \leq \mu \leq \mu'_0 \end{array} \right.$

The values μ_0 and μ'_0 are of no further importance for the performance of the test.

Suppose T is the sequential test of section 2.1 for testing $\mu = \mu_1$ against $\mu = \mu_2$, then this test leads to a decision as soon as

$$(2.2) \quad \frac{\ln B}{\mu_2 - \mu_1} < \sum_{i=1}^n \frac{x_i - \frac{\mu_1 + \mu_2}{2}}{\sigma_i^2} < \frac{\ln A}{\mu_2 - \mu_1}$$

does not hold, where

$$A = \frac{1-\beta}{\alpha}, \quad B = \frac{\beta}{1-\alpha},$$

α = the probability of accepting $\mu \geq \mu_0$ according to T if $\mu = \mu_1$,
 β = the probability of accepting $\mu < \mu_0$ according to T if $\mu = \mu_2$.

Suppose furthermore that T' is the analogous sequential test for testing $\mu = \mu_3$ against $\mu = \mu_4$ then T' leads to a decision as soon as

$$(2.3) \quad \frac{\ln B'}{\mu_4 - \mu_3} < \sum_{i=1}^n \frac{x_i - \frac{\mu_3 + \mu_4}{2}}{\sigma_i^2} < \frac{\ln A'}{\mu_4 - \mu_3}$$

does not hold, where

$$A' = \frac{1-\beta'}{\alpha'}, \quad B' = \frac{\beta'}{1-\alpha'},$$

α' = the probability of accepting $\mu > \mu'_0$ according to T' if $\mu = \mu_3$

β' = the probability of accepting $\mu \leq \mu'_0$ according to T' if $\mu = \mu_4$.

We introduce the following notation

$$(2.4) \quad \left\{ \begin{array}{ll} a = \frac{\ln A}{\mu_2 - \mu_1} & a' = \frac{\ln A'}{\mu_4 - \mu_3} \\ b = \frac{\ln B}{\mu_2 - \mu_1} & b' = \frac{\ln B'}{\mu_4 - \mu_3} \\ \sum_{i=1}^n \frac{x_i - \frac{\mu_1 + \mu_2}{2}}{\sigma_i^2} = y_n & \sum_{i=1}^n \frac{x_i - \frac{\mu_3 + \mu_4}{2}}{\sigma_i^2} = y'_n \end{array} \right.$$

then according to T a decision is taken as soon as

$$(2.2a) \quad b < y_n < a$$

does not hold and according to T' as soon as

$$(2.3a) \quad b' < y'_n < a'$$

does not hold.

From (2.4) it follows that a and a' are positive, b and b' negative and

$$(2.5) \quad y_n > y'_n$$

If now the inequalities

$$(2.6) \quad \begin{cases} b \leq b' \\ a \leq a' \end{cases}$$

are fulfilled, it follows from (2.5) and (2.6) that T' cannot lead to the decision $\mu > \mu_0$ before the decision $\mu \geq \mu_0$ has been found according to T and that T cannot give the decision

$\mu < \mu_0$ before the decision $\mu \leq \mu_0$ has been given by T' .

In that case only the following decision according to T and T' are possible:

1. T' gives the decision $\mu \leq \mu_0$ and T gives (at the same or a later step) the decision $\mu < \mu_0$ or the decision $\mu \geq \mu_0$.
2. T gives the decision $\mu \geq \mu_0$ and T' gives (at the same or a later step) the decision $\mu \leq \mu_0$ or the decision $\mu > \mu_0$.

The sequential test with three possible decisions is then defined as follows:

Additional observations are taken as long as not both tests T and T' have given a decision. As soon as both tests are terminated a decision is taken according to the following rules:

1. $\mu < \mu_0$ if T' has given the decision $\mu \leq \mu_0$ and T the decision $\mu < \mu_0$,
2. $\mu > \mu_0$ if T has given the decision $\mu \geq \mu_0$ and T' the decision $\mu > \mu_0$,
3. $\mu_0 \leq \mu \leq \mu_0$ if T has given the decision $\mu \geq \mu_0$ and T' the decision $\mu \leq \mu_0$.

If (2.6) does not hold there exists the possibility of accepting $\mu > \mu_0$ according to T' and of afterwards accepting $\mu < \mu_0$ ($< \mu_0$) according to T . This kind of contradictory result is excluded by (2.6).

If the random variables x_i all have the same variance σ^2 the test may be carried out graphically, cf. [7].

3. Sequential test with three possible decisions for the comparison of two probabilities.

On the basis of the test of section 2.2 a sequential test with three possible decisions for comparing two unknown probabilities p and p' may be developed as follows.

To the variables a_i and b_i (see section 1) one of the following transformations is applied

$$(3.1) \quad y = 2 \arcsin \sqrt{\frac{n'}{n}} \quad 3)$$

$$(3.2) \quad \begin{cases} y' = 2 \arcsin \sqrt{\frac{n'}{n}} & 0 < n' < n \\ y' = 2 \arcsin \sqrt{\frac{1}{4n}} & n' = 0 \\ y' = \pi - 2 \arcsin \sqrt{\frac{1}{4n}} & n' = n \end{cases} \quad 4)$$

$$(3.3) \quad \begin{cases} y'' = 2 \arcsin \sqrt{\frac{n'}{n}} & 0 < n' < n \\ y'' = \sqrt{\frac{2}{n}} & n' = 0 \\ y'' = \pi - \sqrt{\frac{2}{n}} & n' = n \end{cases}$$

where n' possesses a binomial probability distribution with parameters n and p .

The transformation (3.1) is introduced by FISHER [4], the transformation (3.2) by BARTLETT [1] and (3.3) is given in [8]. For further information about the transformations we refer to [3] (p. 395-416).

Denoting the variables a_i and b_i , after their transformation, by u_i and v_i the sequential test of section 2.2 is applied to the random variables

$$x_i = u_i - v_i \quad (i = 1, 2, 3, \dots)$$

which possess, for large n_i and m_i , approximately a normal probability distribution with mean

$$(3.4) \quad \mu = 2 \arcsin \sqrt{p} - 2 \arcsin \sqrt{p'} = 2 \arcsin (\sqrt{pq} - \sqrt{p'q'}) \quad \begin{matrix} q = 1-p \\ q' = 1-p' \end{matrix}$$

and variance

$$(3.5) \quad \sigma_i^2 = \frac{1}{n_i} + \frac{1}{n_i^2} + \frac{1}{m_i} + \frac{1}{m_i^2}$$

Two values μ_0 and μ'_0 and four values μ_1 , μ_2 , μ_3 and μ_4 of μ must be chosen, with:

$$(3.6) \quad \mu_1 < \mu_0 < \mu_2 < \mu_3 < \mu'_0 < \mu_4$$

3) Tables of $y = 2 \arcsin \sqrt{\alpha}$ are given in [5] for $\alpha = 0,000(0,001)1,000$, p. 70-71, with y in radians.

4) Tables of $y = 2 \arcsin \sqrt{\frac{1}{4n}}$ and $y = \pi - 2 \arcsin \sqrt{\frac{1}{4n}}$ are given in [3], p. 406 for $n = 10(1)50$.

and four values α , α' , β and β' (all $< \frac{1}{2}$) with (see (2.6))

$$(3.7) \quad \begin{cases} \frac{\ln B}{\mu_2 - \mu_1} \leq \frac{\ln B'}{\mu_4 - \mu_3} \\ \frac{\ln A}{\mu_2 - \mu_1} \leq \frac{\ln A'}{\mu_4 - \mu_3} \end{cases},$$

where

$$A = \frac{1-\beta}{\alpha} \quad B = \frac{\beta}{1-\alpha}$$

$$A' = \frac{1-\beta'}{\alpha'} \quad B' = \frac{\beta'}{1-\alpha'}.$$

Having chosen these values the abovementioned test may be applied, leading to one of the decisions:

$$(3.8) \quad \begin{aligned} 1. & \mu < \mu_0 \\ 2. & \mu > \mu'_0 \\ 3. & \mu_0 \leq \mu \leq \mu'_0. \end{aligned}$$

We shall translate these decisions in terms of p and p' .
Let

$$(3.9) \quad \sqrt{pq'} - \sqrt{p'q} = \delta,$$

then

$$(3.10) \quad \mu = 2 \arcsin \delta \quad |\delta| \leq 1.$$

The functional relationship (3.9) between p and p' for given δ^2 consists (see fig. 1) of the arcs PQ and RS of the ellipse:

$$(3.11) \quad p^2 + p'^2 - 2pp'(1-2\delta^2) - 2\delta^2(p+p') + \delta^4 = 0$$

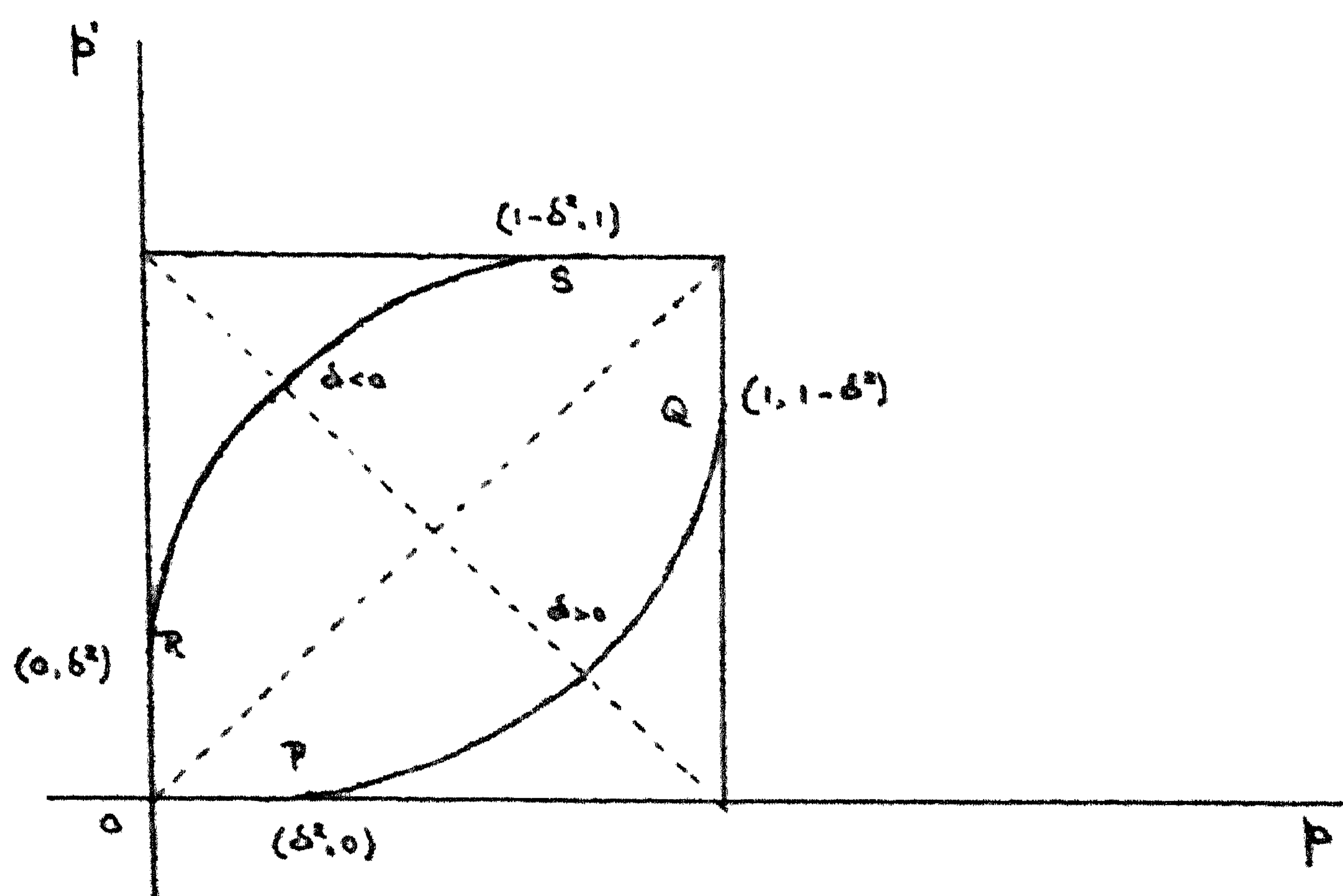


fig. 1.

Functional relationship between p and p' for given value of δ^2 .

Choosing two values δ_0 and δ'_0 and four values δ_1 , δ_2 , δ_3 and δ_4 of δ with

$$(3.12) \quad \delta_1 < \delta_0 < \delta_2 < 0 < \delta_3 < \delta'_0 < \delta_4,$$

the decisions (3.8) are equivalent with the following decisions for δ (see (3.10)):

$$(3.13) \quad \begin{aligned} &1. \delta < \delta_0 \\ &2. \delta > \delta'_0 \\ &3. \delta_0 \leq \delta \leq \delta'_0 \end{aligned}$$

and hence with the following decisions for p and p' :

$$(3.14) \quad \begin{aligned} &1. \text{ the point } (p, p') \text{ lies above the arc RS of figure} \\ &\quad \text{with } \delta = \delta_0, \\ &2. \text{ the point } (p, p') \text{ lies below the arc PQ with } \delta = \delta'_0, \\ &3. \text{ the point } (p, p') \text{ lies on or between the arcs PQ and} \\ &\quad \text{RS with } \delta = \delta'_0 \text{ and } \delta = \delta_0 \text{ respectively.} \end{aligned}$$

The values $\delta_1, \delta_2, \delta_3$ and δ_4 may be chosen by means of fig. 2, where the arcs PQ and RS are given for several values of δ .

One may also choose these values as follows:

1. WALD [6] uses the ratio

$$u = \frac{p q'}{p' q}.$$

On the line $p + p' = 1$ δ can be expressed in terms of u :

$$(3.15) \quad u = \left(\frac{1 + \delta}{1 - \delta} \right)^2,$$

which is equivalent to

$$(3.16) \quad \begin{cases} \delta = 0 & \text{if } u = 1 \\ \delta = \frac{u + 1 - 2\sqrt{u}}{u - 1} & \text{if } u \neq 1. \end{cases}$$

Choosing four values for u with

$$(3.17) \quad u_1 < u_2 < 1 < u_3 < u_4$$

one finds four values for δ such that

$$(3.18) \quad \delta_1 < \delta_2 < 0 < \delta_3 < \delta_4.$$

2. On the line $p + p' = 1$ the equality

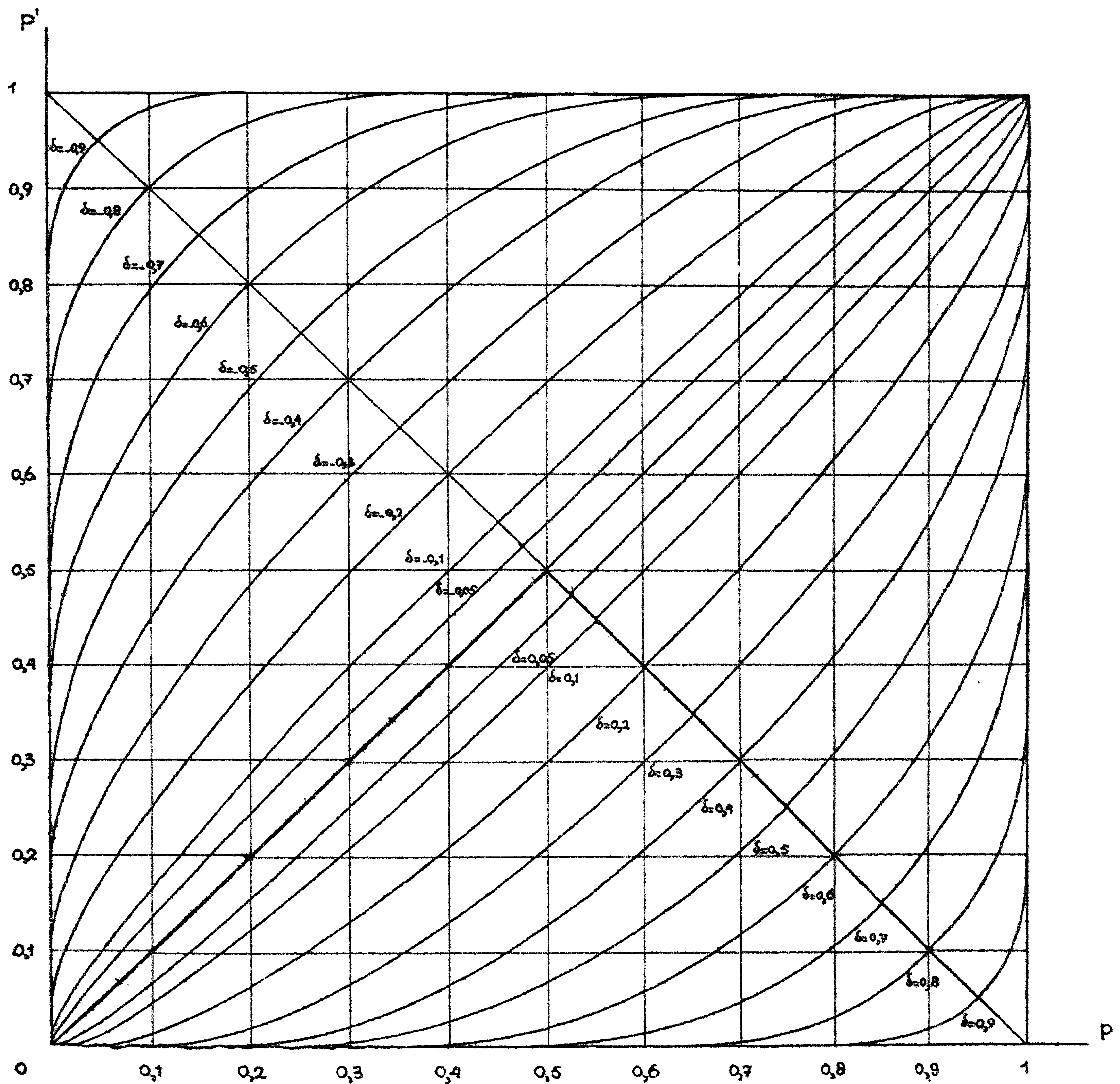
$$(3.19) \quad \delta = p - p'$$

holds.

Choosing four values for $p - p'$ one finds four values for δ . The four values of δ (or u respectively) must furthermore be chosen such that (3.7) holds.

Usually one will choose these values symmetrically, i.e. such that

Figure 2



Functional relationship between p and p' for several values of δ .

$$(3.20) \quad \left\{ \begin{array}{l} \mu_4 = -\mu_1 \\ \mu_3 = -\mu_2 \end{array} \right.$$

which is equivalent to

$$(3.21) \quad \left\{ \begin{array}{l} \delta_4 = -\delta_1 \\ \delta_3 = -\delta_2 \end{array} \right.$$

and to

$$(3.22) \quad u_1 u_4 = u_2 u_3 = 1$$

If (3.20) holds, (3.7) is equivalent to

$$(3.23) \quad \left\{ \begin{array}{l} B \leq B' \\ A \leq A' \end{array} \right.$$

Remark

It is not necessary that p and p' are constants. We only need a constant δ .

References.

- [1] Bartlett, M.S., Some examples of statistical methods of research in agriculture and applied biology, Suppl. Journ. Royal Stat. Soc., 4 (1937), 137-183.
- [2] De Boer, J., Sequential test with three possible decisions for testing an unknown probability, Appl. Sci. Res. 3 (1953), 249-259.
- [3] Eisenhart, C., M.W.Hasteryand, W.A.Wallis, e.a., Selected techniques of statistical analysis for scientific and industrial research and production and management engineering, New York, 1947.
- [4] Fisher, R.A., On the dominance ratio, Proc. Royal Soc. of Edinburgh, 42 (1921-1922).
- [5] Hald, A., Statistical tables and formulas, New York, 1952.
- [6] Wald, A., Sequential analysis, New York, 1947.
- [7] Sobel, M. and A.Wald, A sequential decision procedure for choosing one of three hypotheses concerning the unknown mean of a normal distribution, Ann. Math. Stat., 20 (1949), 502-522.
- [8] Statistical Research Group of the Columbia University, Sequential analysis of statistical data, applications, section 3, New York, 1945.